
Supplementary Material

A Simple Algorithm for Semi-supervised Learning with Improved Generalization Error Bound

Your Name

EMAIL@YOURDOMAIN.EDU

Your Fantastic Institute, 314159 Pi St., Palo Alto, CA 94306 USA

Your CoAuthor's Name

EMAIL@COAUTHORDOMAIN.EDU

Their Fantastic Institute, 27182 Exp St., Toronto, ON M6H 2T1 CANADA

Proof of Proposition 1

First we note that

$$\begin{aligned} \mathbb{E}_{\mathbf{x}} [(\hat{g}(\mathbf{x}) - g_s(\mathbf{x}))^2] &\leq \mathbb{E}_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\gamma_i^* - \alpha_i) \phi_i(\mathbf{x}) \right)^2 \right] \\ &= \mathbb{E}_{\mathbf{x}} \left[\sum_{i=1}^s (\gamma_i^* - \alpha_i)^2 \phi_i^2(\mathbf{x}) \right] + \sum_{i \neq j} (\gamma_i^* - \alpha_i)(\gamma_j^* - \alpha_j) \mathbb{E}_{\mathbf{x}} [\phi_i(\mathbf{x}) \phi_j(\mathbf{x})] \\ &= \|\gamma^* - \alpha^s\|_2^2 \end{aligned}$$

Second, since γ^* is the optimal solution to $\mathcal{L}(\gamma)$, we have

$$\mathcal{L}(\alpha^s) \geq \mathcal{L}(\gamma^*) + (\alpha^s - \gamma^*) Z Z^\top (\alpha^s - \gamma^*)$$

Since $0 \leq \mathcal{L}(\gamma^*) \leq \mathcal{L}(\alpha^s)$, we have

$$(\alpha^s - \gamma^*) Z Z^\top (\alpha^s - \gamma^*) \leq \mathcal{L}(\alpha^s).$$

Then $\|\gamma^* - \alpha^s\|_2^2 \leq \mathcal{L}(\alpha^s) / \lambda_{\min}(Z Z^\top)$. Third, since $\mathcal{L}(\alpha^s)/n$ is the empirical regression error of $g_s(\mathbf{x})$, by the Talagrand inequality (Koltchinskii, 2011) and Lemma 1, we have, with a probability at least $1 - N^{-3}$, $\mathcal{L}(\alpha^s)/n \leq \eta^2$. We complete the proof by combining the above results.

Proof of Proposition 2

To bound $\lambda_{\min}(Z Z^\top)/n$, we need the following proposition.

Proposition 1 (Concentration Inequality). (*Proposition 1 (Smale & Zhou, 2009)*) Let ξ be a random variable on $(\mathcal{X}, P_{\mathcal{X}})$ with values in a Hilbert space $(\mathcal{H}, \|\cdot\|)$. Assume $\|\xi\| \leq M < \infty$ almost sure. Then with a probability at least $1 - \delta$, we have

$$\left\| \frac{1}{n} \sum_{i=1}^m \xi(\mathbf{x}_i) - \mathbb{E}[\xi] \right\| \leq \frac{4M \ln(2/\delta)}{\sqrt{n}}.$$

We rewrite $Z Z^\top / n$ as

$$\frac{1}{n} Z Z^\top = \frac{1}{n} \sum_{i=1}^n \mathbf{z}_i \mathbf{z}_i^\top$$

Let $\|Z\|_2$, $\|Z\|_F$ be the spectral norm and Frobenius norm of Z , respectively. Since

$$\mathbb{E}_i[\mathbf{z}_i \mathbf{z}_i^\top] = I, \|\mathbf{z}_i \mathbf{z}_i^\top\|_F = \mathbf{z}_i^\top \mathbf{z}_i = \sum_{j=1}^s \phi_j^2(\mathbf{x}_i) \leq M(s),$$

using the concentration inequality in above proposition 1 (Smale & Zhou, 2009, Proposition 1), we have, with a probability at least $1 - \delta$,

$$\left\| \frac{1}{n} Z Z^\top - I \right\|_2 \leq \left\| \frac{1}{n} Z Z^\top - I \right\|_F \leq \frac{4M(s) \ln(2/\delta)}{\sqrt{n}}$$

We complete the proof by setting $\delta = N^{-3}$ and using the fact that $\lambda_{\min}(Z Z^\top / n) \geq 1 - \left\| \frac{1}{n} Z Z^\top - I \right\|_2$.

Proof of Lemma 3

We bound $\mathbb{E}_{\mathbf{x}} [(h_s(\mathbf{x}) - f(\mathbf{x}))^2]$ by

$$\mathbb{E}_{\mathbf{x}} [(h_s(\mathbf{x}) - f(\mathbf{x}))^2] \leq 2\mathbb{E}_{\mathbf{x}} [(g_s(\mathbf{x}) - f(\mathbf{x}))^2] + 2\mathbb{E}_{\mathbf{x}} [(h_s(\mathbf{x}) - g_s(\mathbf{x}))^2] \quad (1)$$

Below we will bound the two terms on R.H.S of the above inequality.

To bound the first term in (1), we use Proposition 1, and with a probability $1 - 2N^{-3}$, we have

$$\|L - \widehat{L}_N\|_{HS} \leq \frac{12 \ln N}{\sqrt{N}} = \tau_N$$

According to Lidskii's inequality, we have

$$\sum_i |\lambda_i - \widehat{\lambda}_i| \leq \frac{12 \ln N}{\sqrt{N}} = \tau_N$$

Following the same analysis as Lemma 1, we have

$$\begin{aligned} \sum_{i=s+1}^{\infty} \alpha_i^2 &\leq R^2 \sum_{i=s+1}^{\infty} \lambda_i \leq R^2 \sum_{i=s+1}^N \widehat{\lambda}_i + R^2 \sum_i |\lambda_i - \widehat{\lambda}_i| \\ &\leq R^2 \left(\frac{a^2}{s^{p-1}} + \frac{12 \ln N}{\sqrt{N}} \right) \leq \frac{2R^2 a^2}{s^{p-1}} \end{aligned}$$

where the last step we use the condition

$$\frac{12 \ln N}{\sqrt{N}} \leq \frac{a^2}{s^{p-1}}$$

Hence, by the same analysis in the proof of Lemma 1, with a probability $1 - 2N^{-3}$, we have

$$\mathbb{E}_{\mathbf{x}} [(f(\mathbf{x}) - g_s(\mathbf{x}))^2] \leq \frac{4a^2 R^2}{s^{p-1}} + 2\epsilon^2 \leq 2\epsilon_s^2$$

To bound the second term on (1), we use the following corollary.

Corollary 1. *Let N be a sufficiently large number such that $\widehat{\phi}_i \in \text{span}(\phi_1, \dots, \phi_N)$. Define*

$$\begin{aligned} \Theta &= (\widehat{\phi}_1, \dots, \widehat{\phi}_s), \\ \Phi &= (\sqrt{\lambda_1} \phi_s, \dots, \sqrt{\lambda_s} \phi_s), \\ \overline{\Phi} &= (\sqrt{\lambda_{s+1}} \phi_{s+1}, \dots, \sqrt{\lambda_N} \phi_N) \end{aligned}$$

Assume

$$r_s = (\lambda_s - \lambda_{s+1}) > 3\|L - \widehat{L}_N\|_{HS}.$$

Then, there exists a matrix $P \in \mathbb{R}^{(N-s) \times s}$ satisfying

$$\|P\|_F \leq \frac{3\|L - \widehat{L}_N\|_{HS}}{r_s}$$

such that

$$\Theta = (\Phi + \overline{\Phi}P)(I + P^\top P)^{-1/2}$$

The above lemma follows from the following perturbation result.

Corollary 2. (Theorem 2.7 of Chapter 6 (Stewart & Guang Sun, 1990)) Let $(\lambda_i, \mathbf{v}_i), i \in [n]$ be the eigenvalues and eigenvectors of a symmetric matrix $A \in \mathbb{R}^{n \times n}$ ranked in the descending order of eigenvalues. Set $X = (\mathbf{v}_1, \dots, \mathbf{v}_r)$ and $Y = (\mathbf{v}_{r+1}, \dots, \mathbf{v}_n)$. Given a symmetric perturbation matrix E , let

$$\widehat{E} = (X, Y)^\top E(X, Y) = \begin{pmatrix} \widehat{E}_{11} & \widehat{E}_{12} \\ \widehat{E}_{21} & \widehat{E}_{22} \end{pmatrix}$$

Let $\|\cdot\|$ represent a consistent family of norms and set

$$\gamma = \|\widehat{E}_{21}\|, \delta = \lambda_r - \lambda_{r+1} - \|\widehat{E}_{11}\| - \|\widehat{E}_{22}\|$$

If $\delta > 0$ and $2\gamma < \delta$, then there exists a unique matrix $P \in \mathbb{R}^{(n-r) \times r}$ satisfying $\|P\| < \frac{2\gamma}{\delta}$ such that

$$\begin{aligned} X' &= (X + YP)(I + P^\top P)^{-1/2} \\ Y' &= (Y - XP^\top)(I + PP^\top)^{-1/2} \end{aligned}$$

are the eigenvectors of $A + E$.

Proof of Corollary 1. Let $\varphi_i = \sqrt{\lambda_i}\phi_i$, it can be shown that $\langle \varphi_i, \varphi_j \rangle_{\mathcal{H}_\kappa} = \delta_{ij}$. Define matrix B as

$$B_{i,j} = \frac{1}{N} \sum_{k=1}^N \widehat{\lambda}_k \langle \widehat{\phi}_k, \varphi_i \rangle \langle \widehat{\phi}_k, \varphi_j \rangle.$$

Let \mathbf{z}_i be the eigenvector of B corresponding to eigenvalue $\widehat{\lambda}_i/N$. It is straightforward to show that

$$\mathbf{z}_i = (\langle \varphi_1, \widehat{\phi}_i \rangle_{\mathcal{H}_\kappa}, \dots, \langle \varphi_N, \widehat{\phi}_i \rangle_{\mathcal{H}_\kappa})^\top, i \in [N]$$

and therefore we have

$$\widehat{\phi}_i = \sum_{k=1}^N z_{i,k} \varphi_k, i \in [N], \text{ or } \Theta = (\Phi, \overline{\Phi})Z$$

where $Z = (\mathbf{z}_1, \dots, \mathbf{z}_s)$. To decide the relationship between $\{\widehat{\phi}_i\}_{i=1}^s$ and $\{\varphi_i\}_{i=1}^N$, we need to determine matrix Z . We define matrix $D = \text{diag}(\lambda_1, \dots, \lambda_N)$ and matrix $E = B - D$, i.e.

$$E_{i,j} = B_{i,j} - \lambda_i \delta_{i,j} = \langle \varphi_i, (\widehat{L}_N - L)\varphi_j \rangle_{\mathcal{H}_\kappa}.$$

Following the notation of Theorem 2, we define $X = (e_1, \dots, e_s)$ and $Y = (e_{s+1}, \dots, e_N)$, where e_1, \dots, e_N are the canonical bases of \mathbb{R}^N , which are also eigenvectors of D . Define δ and γ as follows

$$\begin{aligned} \gamma &= \sqrt{\sum_{i=1}^s \sum_{j=s+1}^N \langle \varphi_i, (L - \widehat{L}_N)\varphi_j \rangle_{\mathcal{H}_\kappa}^2} \\ \delta &= r_s - \sqrt{\sum_{i,j=1}^s \langle \varphi_i, (L - \widehat{L}_N)\varphi_j \rangle_{\mathcal{H}_\kappa}^2} - \sqrt{\sum_{i,j=s+1}^N \langle \varphi_i, (L - \widehat{L}_N)\varphi_j \rangle_{\mathcal{H}_\kappa}^2} \end{aligned}$$

where $r_s = \lambda_s - \lambda_{s+1}$. It is easy to verify that γ, δ are defined with respect to the Frobenius norm of \widehat{E} in Theorem 2. In order to apply the result in Theorem 2, we need to show $\delta > 0$ and $\gamma < \delta/2$. To this end, we need to provide the lower and upper bounds for γ and δ , respectively. We first bound δ as

$$\delta - r_s \geq -\sqrt{\sum_{i,j=1}^N \langle \varphi_i, (L - \widehat{L}_N) \varphi_j \rangle_{\mathcal{H}_\kappa}^2} = -\|L - \widehat{L}_N\|_{HS}$$

We then bound γ as

$$\gamma = \sqrt{\sum_{i=1}^r \sum_{j=r+1}^N \langle \varphi_i, (L - \widehat{L}_N) \varphi_j \rangle_{\mathcal{H}_\kappa}^2} \leq \sqrt{\sum_{i=1}^N \sum_{j=1}^N \langle \varphi_i, (L - \widehat{L}_N) \varphi_j \rangle_{\mathcal{H}_\kappa}^2} = \|L - \widehat{L}_N\|_{HS}$$

Hence, when $r_s > 3\|L - \widehat{L}_N\|_{HS}$, we have $\delta > 2\gamma > 0$, which satisfies the condition specified in Theorem 2. Thus, according to Theorem 2, there exists a $P \in \mathbb{R}^{(N-s) \times s}$ satisfying $\|P\| < 2\gamma/\delta$, such that

$$Z = (\mathbf{z}_1, \dots, \mathbf{z}_s) = (X + YP)(I + P^\top P)^{-1/2}$$

implying

$$\Theta = (\Phi, \overline{\Phi})Z = (\Phi + \overline{\Phi}P)(I + P^\top P)^{-1/2}$$

□

By Corollary 1, since $r_s \geq 3\tau_N^{2/3} > 3\tau_N \geq 3\|L - \widehat{L}_N\|_{HS}$, by the above theorem, we have

$$\begin{aligned} \sum_{i=1}^s \|\widehat{\phi}_i - \sqrt{\lambda_i} \phi_i\|_{\mathcal{H}_\kappa}^2 &= \|\Theta - \Phi\|_F^2 = \|\Phi(I - [I + P^\top P]^{-1/2})\|_F^2 + \|\overline{\Phi}P(I + P^\top P)^{-1/2}\|_F^2 \\ &\leq 2\|P^\top P\|_F^2 \leq \frac{18\|L - \widehat{L}_N\|_{HS}^2}{r_s^2} \leq \frac{18\tau_N^2}{r_s^2} \quad (\text{w.p. } 1 - 2N^{-3}) \end{aligned}$$

Then, we have

$$\begin{aligned} \mathbb{E}_{\mathbf{x}} [(h_s(\mathbf{x}) - g_s(\mathbf{x}))^2] &= \mathbb{E}_{\mathbf{x}} \left[\left(\sum_{i=1}^s \frac{\alpha_i}{\sqrt{\lambda_i}} (\widehat{\phi}_i(\mathbf{x}) - \sqrt{\lambda_i} \phi_i(\mathbf{x})) \right)^2 \right] \\ &= \sum_{i=1}^s \frac{\alpha_i^2}{\lambda_i} \mathbb{E}_{\mathbf{x}} \left[\sum_{i=1}^s (\widehat{\phi}_i(\mathbf{x}) - \sqrt{\lambda_i} \phi_i(\mathbf{x}))^2 \right] \leq \sum_{i=1}^s \frac{\alpha_i^2}{\lambda_i} \sum_{i=1}^s \|\widehat{\phi}_i(\cdot) - \sqrt{\lambda_i} \phi_i(\cdot)\|_{\mathcal{H}_\kappa}^2 \leq \frac{18\tau_N^2 R^2}{r_s^2} \quad (\text{w.p. } 1 - 2N^{-3}) \end{aligned}$$

Combining the above results, with a probability $1 - 2N^{-3}$, we have

$$\mathbb{E}_{\mathbf{x}} [(h_s(\mathbf{x}) - f(\mathbf{x}))^2] \leq 4\epsilon_s^2 + \frac{36\tau_N^2 R^2}{Nr_s^2}$$

1. Proof of Lemma 4

We bound as follows:

$$\begin{aligned} \mathbb{E}_{\mathbf{x}} [(\widehat{g}(\mathbf{x}) - h_s(\mathbf{x}))^2] &\leq \mathbb{E}_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\widehat{\gamma}_i^* - \alpha_i) \frac{\widehat{\phi}_i(\mathbf{x})}{\sqrt{\lambda_i}} \right)^2 \right] \\ &= \mathbb{E}_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\widehat{\gamma}_i^* - \alpha_i) \phi_i(\mathbf{x}) + (\widehat{\gamma}_i^* - \alpha_i) \left(\frac{\widehat{\phi}_i(\mathbf{x})}{\sqrt{\lambda_i}} - \phi_i(\mathbf{x}) \right) \right)^2 \right] \\ &\leq 2\mathbb{E}_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\widehat{\gamma}_i^* - \alpha_i) \phi_i(\mathbf{x}) \right)^2 \right] + 2\mathbb{E}_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\widehat{\gamma}_i^* - \alpha_i) \left(\frac{\widehat{\phi}_i(\mathbf{x})}{\sqrt{\lambda_i}} - \phi_i(\mathbf{x}) \right) \right)^2 \right] \end{aligned}$$

For the first term in the above inequality, using the fact $E_{\mathbf{x}}[\phi_i(\mathbf{x})\phi_j(\mathbf{x})] = \delta_{ij}$, we have

$$E_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\hat{\gamma}_i^* - \alpha_i) \phi_i(\mathbf{x}) \right)^2 \right] = \|\gamma^* - \alpha^s\|_2^2$$

For the second term, we bound it as

$$\begin{aligned} E_{\mathbf{x}} \left[\left(\sum_{i=1}^s (\hat{\gamma}_i^* - \alpha_i) \left(\frac{\hat{\phi}_i(\mathbf{x})}{\sqrt{\lambda_i}} - \phi_i(\mathbf{x}) \right) \right)^2 \right] &\leq \|\hat{\gamma}^* - \alpha^s\|_2^2 E_{\mathbf{x}} \left[\sum_{i=1}^s \left(\frac{\hat{\phi}_i(\mathbf{x})}{\sqrt{\lambda_i}} - \phi_i(\mathbf{x}) \right)^2 \right] \\ &\leq \frac{\|\hat{\gamma}^* - \alpha^s\|_2^2}{\lambda_s} E_{\mathbf{x}} \left[\sum_{i=1}^s \left(\hat{\phi}_i(\mathbf{x}) - \sqrt{\lambda_i} \phi_i(\mathbf{x}) \right)^2 \right] \leq \frac{18\tau_N^2 \|\hat{\gamma}^* - \alpha^s\|_2^2}{\lambda_s r_s^2} \leq \frac{18\tau_N^2 \|\hat{\gamma}^* - \alpha^s\|_2^2}{r_s^3} \quad (\text{w.p. } 1 - 2N^{-3}) \end{aligned}$$

Similar to the infinite case, we introduce $\mathbf{z}_i = (\hat{\phi}_1(\mathbf{x}_i)/\sqrt{\lambda_1}, \dots, \hat{\phi}_s(\mathbf{x}_i)/\sqrt{\lambda_s})^\top$ and $Z = (\mathbf{z}_1, \dots, \mathbf{z}_n)$. Then by the similar analysis to Proposition 1 and Proposition 2, with a probability $1 - 2N^{-3}$, we have $\|\alpha^s - \hat{\gamma}^*\|_2 \leq n\hat{\eta}^2/\lambda_{\min}(ZZ^\top) \leq 2\hat{\eta}^2$. We then complete the proof by using the assumption **B3** that $r_s^3 \geq 27\tau_N^2$.

References

- Koltchinskii, Vladimir. *Oracle Inequalities in Empirical Risk Minimization and Sparse Recovery Problems*. Springer, 2011.
- Smale, Steve and Zhou, Ding-Xuan. Geometry on probability spaces. *Constructive Approximation*, 30:311–323, 2009.
- Stewart, G. W. and Guang Sun, Ji. *Matrix Perturbation Theory*. Academic Press, 1990.